



TECNICAS DE PRONOSTICOS

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CONTENIDO

VII. FUNCIONES ORTOGONALES EMPÍRICAS/ANÁLISIS DE
COMPONENTES PRINCIPALES

COVARIANZA

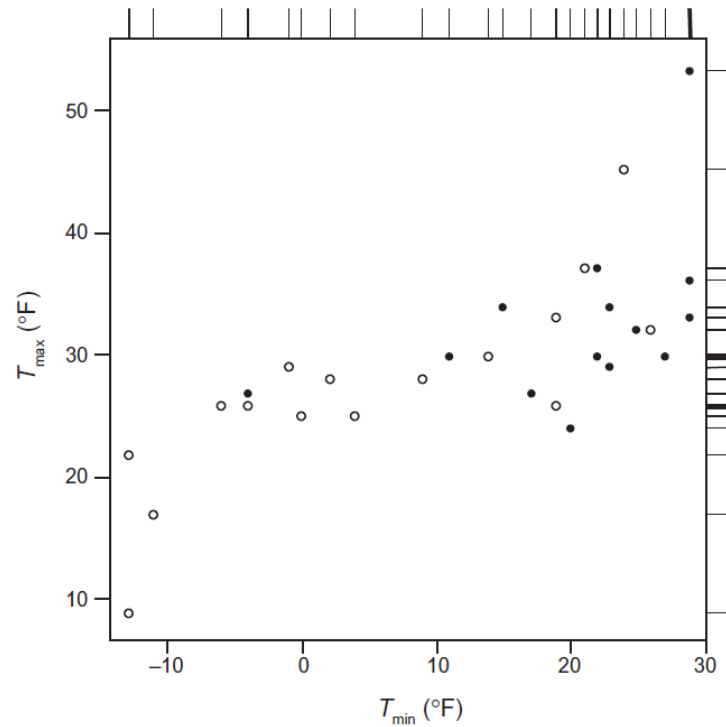


FIGURE 3.17 Scatterplot for daily maximum and minimum temperatures during January 1987 at Ithaca, New York. “Fringes” along the margins separately indicate the individual data distributions, with repeated data represented by heavier lines. Closed circles represent days with at least 0.01 in. of precipitation (liquid equivalent).

$$S_{X,Y} = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{(n - 1)}$$

CORRELACION

Correlación de Pearson

Una forma de ver la correlación de Pearson es como la relación entre la covarianza muestral entre las dos variables, y el producto de las dos desviaciones estándar; donde los numeradores denotan anomalías, o resta de valores medios.

$$r_{x,y} = \frac{\text{Cov}(x, y)}{s_x s_y} = \frac{\frac{1}{n-1} \sum_{i=1}^n [(x_i - \bar{x})(y_i - \bar{y})]}{\left[\frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2 \right]^{1/2} \left[\frac{1}{n-1} \sum_{i=1}^n (y_i - \bar{y})^2 \right]^{1/2}},$$

$$= \frac{\sum_{i=1}^n (x'_i y'_i)}{\left[\sum_{i=1}^n (x'_i)^2 \right]^{1/2} \left[\sum_{i=1}^n (y'_i)^2 \right]^{1/2}}$$

Varianza (S^2)

$$S^2 = \frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n-1}$$

Desviación estándar (S)

$$S = \sqrt{\frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n-1}}$$

$$r_{x,y} = \frac{1}{n-1} \sum_{i=1}^n \left[\frac{(x_i - \bar{x})}{s_x} \frac{(y_i - \bar{y})}{s_y} \right] = \frac{1}{n-1} \sum_{i=1}^n z_{x_i} z_{y_i},$$

ANALISIS DE FUNCIONES ORTOGONALES EMPÍRICAS / COMPONENTES PRINCIPALES

Statistical Methods for Signal Detection in Climate

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DCESS Report # 2

January, 2001

PREPARACION DE DATOS

Time $\longrightarrow$

$$\mathbf{F} = \begin{bmatrix} F_1(1) & F_1(2) & \dots & F_1(N) \\ F_2(1) & F_2(2) & \dots & F_2(N) \\ \dots & \dots & \dots & \dots \\ F_M(1) & F_M(2) & \dots & F_M(N) \end{bmatrix} \quad \downarrow \text{Location}$$

Matriz de F de anomalías

$$F_m(t) = \frac{\psi_m(t) - \mu_m}{\sigma_m} \quad (2.1)$$

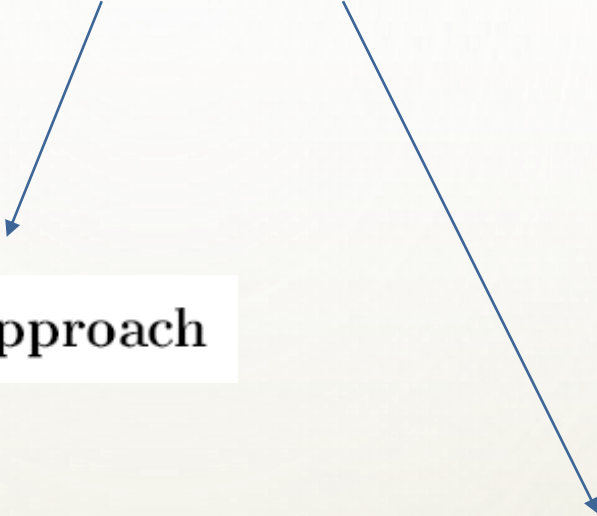
where μ_m is the record mean:

$$\mu_m = \frac{1}{N} \sum_{t=1}^N \psi_m(t) \quad (2.2)$$

and σ_m is the record standard deviation:

$$\sigma_m = \left[\frac{1}{N-1} \sum_{t=1}^N \psi_m^2(t) \right]^{1/2} \quad (2.3)$$

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The Covariance Matrix Approach

The Singular Value Decomposition Approach

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The Covariance Matrix Approach

Matriz de Covarianza $\longrightarrow \mathbf{R}_{\mathbf{FF}} = \mathbf{F} * \mathbf{F}^\dagger$

$$\mathbf{R}_{\mathbf{FF}} = \begin{bmatrix} \langle F_1 F_1 \rangle & \langle F_1 F_2 \rangle & \dots & \langle F_1 F_M \rangle \\ \langle F_2 F_1 \rangle & \langle F_2 F_2 \rangle & \dots & \langle F_2 F_M \rangle \\ \dots & \dots & \dots & \dots \\ \langle F_M F_1 \rangle & \langle F_M F_2 \rangle & \dots & \langle F_M F_M \rangle \end{bmatrix}$$

$$\langle F_i F_j \rangle = \langle F_j F_i \rangle = \frac{1}{N-1} \sum_{t=1}^N F_i(t) F_j(t)$$

ANÁLISIS DE FUNCIONES ORTOGONALES EMPÍRICAS / COMPONENTES PRINCIPALES

The Covariance Matrix Approach

$$\mathbf{R}_{FF} * \mathbf{E} = \mathbf{E} * \mathbf{\Lambda}$$

$M \times M \quad M \times M$

Eigen Value Problem

loadings

$$\mathbf{E} = \begin{bmatrix} E_1^1 & E_1^2 & \dots & E_1^M \\ E_2^1 & E_2^2 & \dots & E_2^M \\ \vdots & \vdots & \ddots & \vdots \\ E_M^1 & E_M^2 & \dots & E_M^M \end{bmatrix}$$

$\downarrow \quad \downarrow \quad \quad \downarrow$
 $E^1 \quad E^2 \quad \quad E^M \longrightarrow \text{Eigenvectors } E^k$

Eigenvectors

$$\mathbf{\Lambda} = \begin{bmatrix} \lambda_1 & 0 & \dots & 0 \\ 0 & \lambda_2 & \dots & 0 \\ \dots & \dots & \ddots & \dots \\ 0 & 0 & \dots & \lambda_M \end{bmatrix}$$

Eigenvalues

Componentes
Principales

$$A^k(t) = \sum_{m=1}^M E_m^k F_m(t)$$

$$\mathbf{A} = \mathbf{E}^\dagger * \mathbf{F}$$

$M \times N$

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The Covariance Matrix Approach

$$\% \text{ Variance Mode } k = \frac{\lambda_k}{\sum_{i=1}^K \lambda_i} * 100$$

The original field F can be totally reconstructed by multiplying each EOF pattern E^k by its corresponding principal component A^k and adding the products over all K modes:

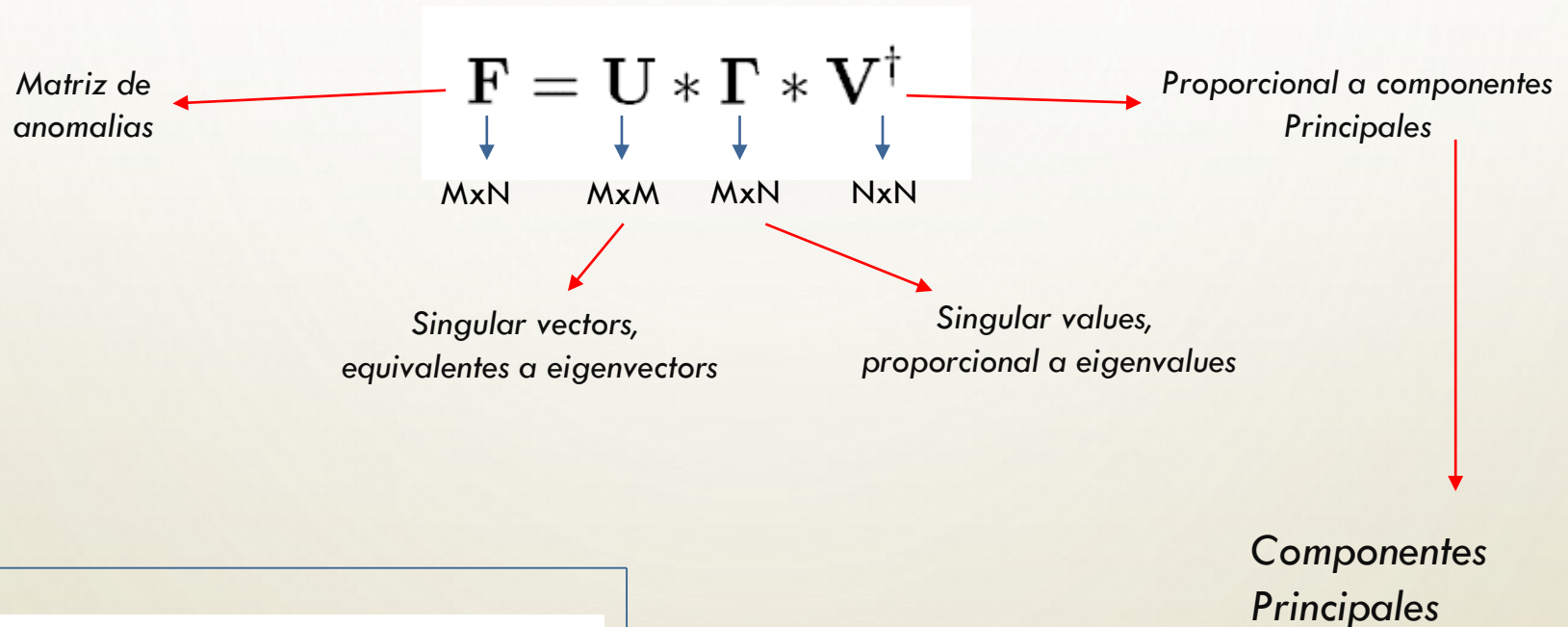
$$F_m(t) = \sum_{k=1}^K E_m^k A^k(t)$$

In matrix notation:

$$\mathbf{F} = \mathbf{E} * \mathbf{A}$$

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The Singular Value Decomposition Approach



we can reconstruct field F adding all K modes of the decomposition as:

$$F_m(t) = \sum_{k=1}^K U_m^k \gamma_k V^{\dagger k}(t)$$

$$\mathbf{A} = \mathbf{\Gamma} * \mathbf{V}^\dagger$$

$$A^k(t) = \gamma_k V^{\dagger k}(t)$$

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UNIDADES Y PRESENTACION

$$A^{k*}(t) = \frac{A^k(t)}{\sqrt{\lambda_k}}$$

$$E_m^{k*} = E_m^k \sqrt{\lambda_k}$$

$$\mathbf{r}_m^k = [A^k(t), F_m(t)]$$

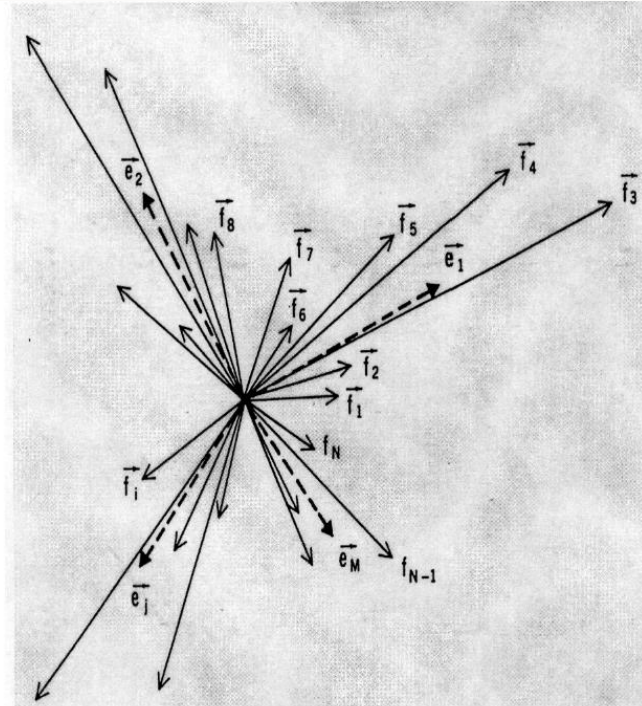


Figure 2.1: Example of a possible configuration of the data vectors f_n ($n = 1 \dots N$ denote the time steps) and the empirical orthogonal vectors e_m , $m = 1 \dots M$. From Peixoto and Oort (1992)

PRACTICA

ANALISIS DE FUNCIONES ORTOGONALES EMPÍRICAS / COMPONENTES PRINCIPALES

Combined Empirical Orthogonal Functions

$$\mathbf{F} = \begin{bmatrix} S_1(1) & S_1(2) & \dots & S_1(N) \\ S_2(1) & S_2(2) & \dots & S_2(N) \\ \dots & \dots & \dots & \dots \\ S_{M_s}(1) & S_{M_s}(2) & \dots & S_{M_s}(N) \\ P_1(1) & P_1(2) & \dots & P_1(N) \\ P_2(1) & P_2(2) & \dots & P_2(N) \\ \dots & \dots & \dots & \dots \\ P_{M_p}(1) & P_{M_p}(2) & \dots & P_{M_p}(N) \end{bmatrix}$$

$$\mathbf{R}_{\mathbf{FF}} = \begin{bmatrix} \langle S_1 S_1 \rangle & \dots & \langle S_1 S_{M_s} \rangle & \langle S_1 P_1 \rangle & \dots & \langle S_1 P_{M_p} \rangle \\ \dots & \dots & \dots & \dots & \dots & \dots \\ \langle S_{M_s} S_1 \rangle & \dots & \langle S_{M_s} S_{M_s} \rangle & \langle S_{M_s} P_1 \rangle & \dots & \langle S_{M_s} P_{M_p} \rangle \\ \langle P_1 S_1 \rangle & \dots & \langle P_1 S_{M_s} \rangle & \langle P_1 P_1 \rangle & \dots & \langle P_1 P_{M_p} \rangle \\ \dots & \dots & \dots & \dots & \dots & \dots \\ \langle P_{M_p} S_1 \rangle & \dots & \langle P_{M_p} S_{M_s} \rangle & \langle P_{M_p} P_1 \rangle & \dots & \langle P_{M_p} P_{M_p} \rangle \end{bmatrix}$$

$$\mathbf{E} * \mathbf{R}_{\mathbf{FF}} = \mathbf{\Lambda} * \mathbf{E}$$

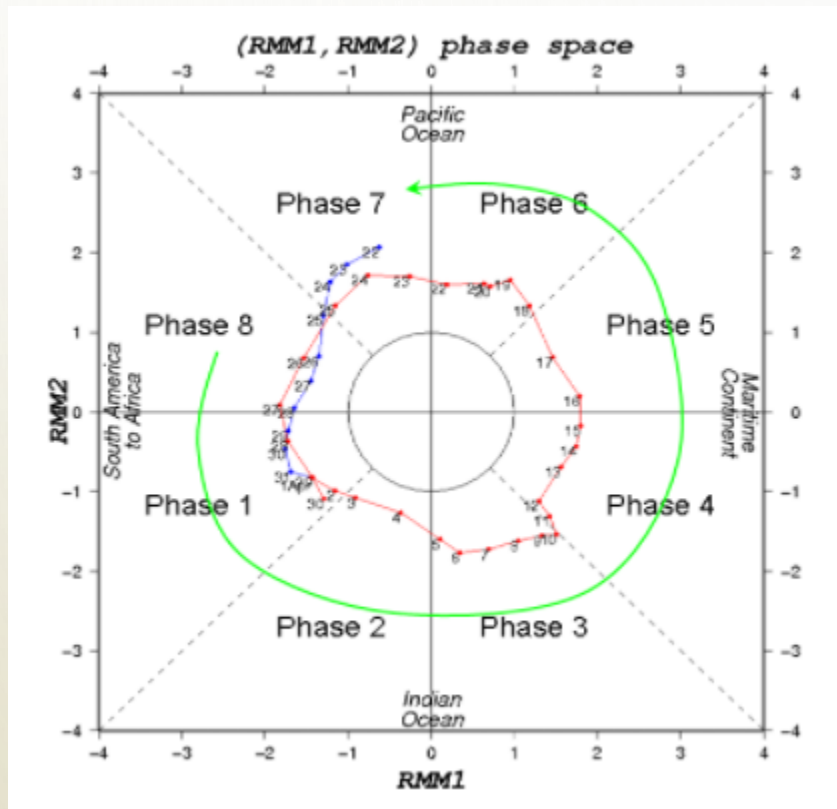
$$\mathbf{R}_{\mathbf{SP}} = \begin{bmatrix} \langle S_1 P_1 \rangle & \langle S_1 P_2 \rangle & \dots & \langle S_1 P_{M_p} \rangle \\ \langle S_2 P_1 \rangle & \langle S_2 P_2 \rangle & \dots & \langle S_2 P_{M_p} \rangle \\ \dots & \dots & \dots & \dots \\ \langle S_{M_s} P_1 \rangle & \langle S_{M_s} P_2 \rangle & \dots & \langle S_{M_s} P_{M_p} \rangle \end{bmatrix}$$

$$\mathbf{R}_{\mathbf{SP}} = \mathbf{U} * \mathbf{\Gamma} * \mathbf{V}^\dagger$$

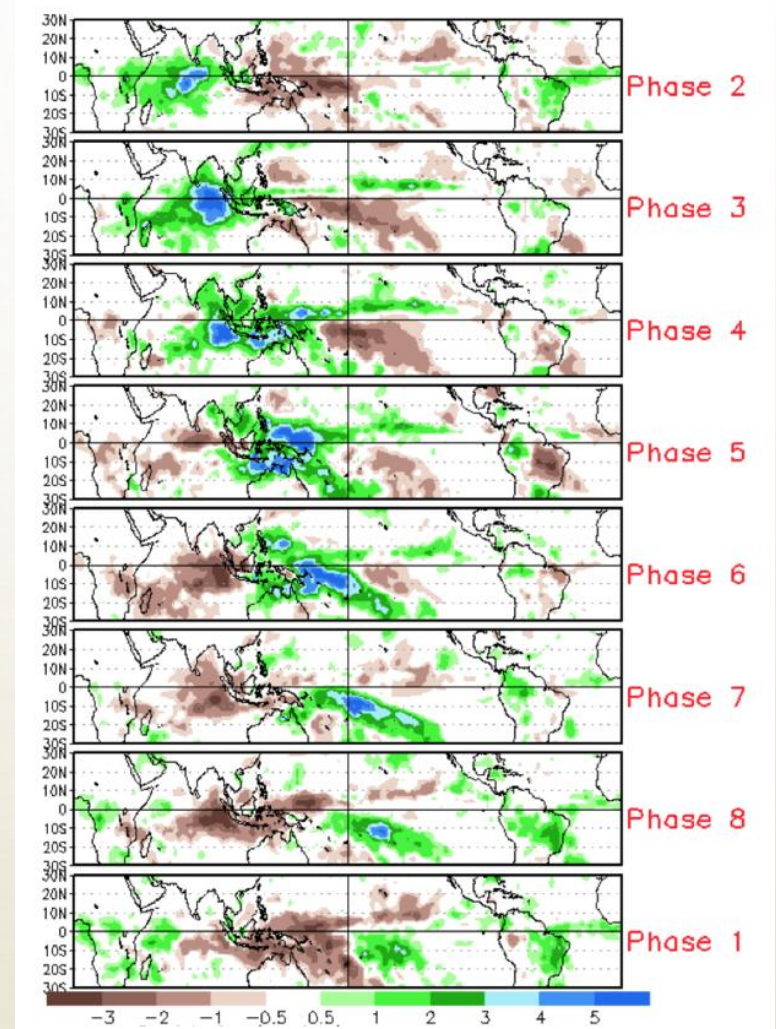
Combined Empirical Orthogonal Functions

EJEMPLO: MJO INDEX

Campos: WIND850 , WIND200, OLR



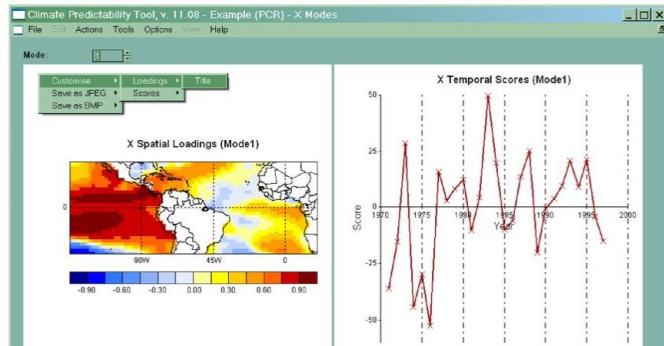
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<https://www.climate.gov/news-features/blogs/enso/what-mjo-and-why-do-we-care>

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OTRAS APLICACIONES



https://iri.columbia.edu/wp-content/uploads/2013/07/CPT_Tutorial.pdf

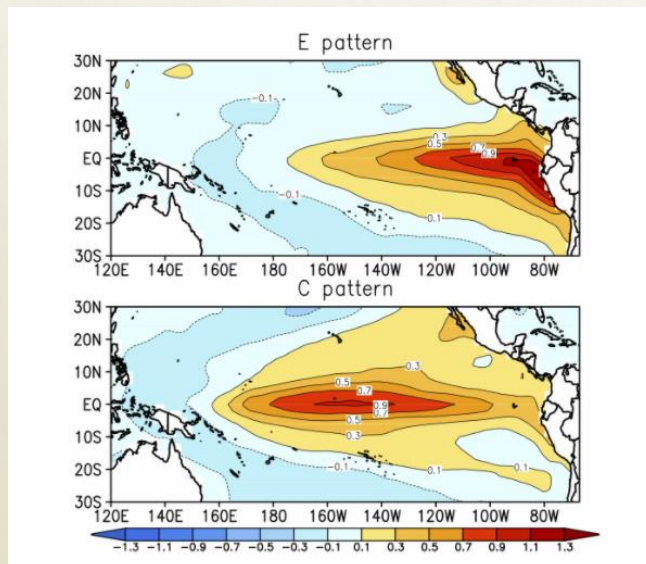
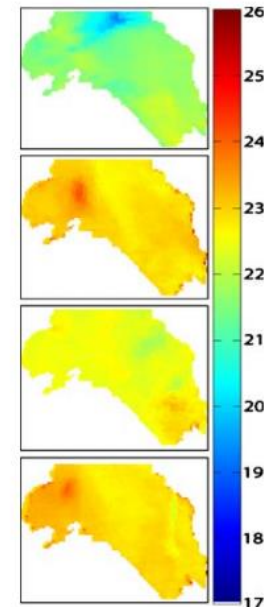
An improved methodology for filling missing values in spatiotemporal climate data set

Application to Tanganyika Lake data set

Antti Sorjamaa · Amaury Lendasse ·
Yves Cornet · Eric Deleersnijder

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Fig. 7 Southern part of the Tanganyika Lake data set. The pictures are the same examples filled by the EOF pruning as in Fig. 3



http://met.igp.gob.pe/el_nino/indices.html